

Regular and anomalous scaling of a randomly advected passive scalar

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Extending Kolmogorov's refined similarity hypothesis to study the inertial behavior $\langle [T(\mathbf{x}+\mathbf{r},t) - T(\mathbf{x},t)]^{2n} \rangle \propto r^{\zeta_{2n}}$ of a passive scalar $T(\mathbf{x},t)$ advected by a rapidly changing incompressible velocity field, a random variable θ was introduced by Ching [Phys. Rev. Lett. **79**, 3644 (1997)]. In this paper, the statistical distribution of the random variable $X = \theta / \sqrt{\langle \theta^2 \rangle}$ is investigated analytically for the scaling in two limits, n -independent scaling $\zeta_{2n} = \zeta_2$ and regular scaling $\zeta_{2n} = n\zeta_2$, and numerically for the scaling of the Kraichnan conjecture $\zeta_{2n} = \frac{1}{2}[\sqrt{4nd\zeta_2 + (d - \zeta_2)^2} - (d - \zeta_2)]$. For n -independent scaling $\zeta_{2n} = \zeta_2$, the statistical distribution of X tends to an exponential distribution when $\zeta_2 \rightarrow 0$ or $d \rightarrow \infty$ and to a Gaussian distribution when $\zeta_2 \rightarrow 2$ and $d = 2$. For regular scaling $\zeta_{2n} = n\zeta_2$, the statistical distribution of X tends to a Gaussian distribution when $\zeta_2 \rightarrow 0$ or $d \rightarrow \infty$. In $d = 2$, there seems to be a phase transition for the probability density function $P(X)$ from a convex to a concave function when the value of ζ_2 is increased and the critical point is $\zeta_2 = 4/3$ where the random variable X has a uniform distribution in $[-\sqrt{3}, \sqrt{3}]$. In $d = 3$, $P(X)$ is a convex function for all $0 < \zeta_2 < 2$ and tends to a constant on its support $[-\sqrt{3}, \sqrt{3}]$ when $\zeta_2 \rightarrow 2$. For the scaling of the Kraichnan conjecture, $P(X)$ has two peaks in $d = 2$ for $\zeta_2 > 1.33$, but, in $d = 3$, it has only one peak for all $0 < \zeta_2 < 2$ and changes very slowly with the value of X in the neighborhood of $X = 0$ as $\zeta_2 \rightarrow 2$.

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The problem of passive scalar advection is of physical importance in itself and may also serve as starting point in studying multiscaling in turbulence. Although the governing equation for a passive scalar is apparently linear when a random velocity field is prescribed, the statistical properties of the passive scalar are much more elusive. Nevertheless, a mathematically tractable model in which the velocity field is rapidly changing in time was introduced by Kraichnan [1]. Some results of this model can be obtained exactly. Based upon a linear ansatz [2], Kraichnan gave a conjecture for the scaling exponents ζ_{2n} of the scalar structure functions $S_{2n}(r)$ defined by $S_{2n}(r) \sim r^{\zeta_{2n}}$, where $S_{2n}(r) \equiv \langle [T(\mathbf{x}+\mathbf{r},t) - T(\mathbf{x},t)]^{2n} \rangle \equiv \langle [T_r(\mathbf{x})]^{2n} \rangle$. The simplified model has attracted much attention as a possible model to study intermittency and multiscaling in turbulence [2–9]. Kolmogorov's refined similarity hypothesis (RSH) [10,11] was extended to study the intermittency of this model by Ching [12]. With the RSH formulated in terms of a random variable θ , the molecular-diffusion terms were evaluated. The scaling exponents ζ_{2n} of the scalar structure functions $S_{2n}(r)$ are determined solely by the statistics of the random variable θ . Therefore, the relation between the scaling exponents ζ_{2n} and the statistical distribution of θ is of great interest.

The Hölder reliability inequalities require that $d\zeta_{2n}/dn$ be a nonincreasing function of n . This implies that $\zeta_{2n} \leq n\zeta_2$. The limit $\zeta_{2n} = n\zeta_2$ is called regular scaling. Away from this limit, the scaling is called anomalous scaling. On the other hand, all the theoretical analysis and numerical simulations suggest that ζ_{2n} is a nondecreasing function of n [2,3,8,9]. This requires $\zeta_{2n} \geq \zeta_2$. The physical picture for the limit $\zeta_{2n} = \zeta_2$ was given in [2]. The Kraichnan conjecture based upon the linear ansatz suggests that the scaling is anomalous midway between the two limits $\zeta_{2n} = \zeta_2$ and $\zeta_{2n} = n\zeta_2$. The scaling exponent ζ_2 is taken to be in the region $0 < \zeta_2 < 2$ [2]. In this paper, we investigate the statistical

properties of θ for the scaling in the two limits of n -independent scaling $\zeta_{2n} = \zeta_2$ and regular scaling $\zeta_{2n} = n\zeta_2$, and also for the scaling of Kraichnan conjecture.

Consider the structure functions $S_{2n}(r) = \langle [T_r(\mathbf{x})]^{2n} \rangle$ of a passive scalar field $T(\mathbf{x},t)$ that obeys

$$\frac{\partial T}{\partial t} + (\mathbf{v} \cdot \nabla)T = \kappa \nabla^2 T. \quad (1)$$

In the limit of an infinitely rapid change in time of $\mathbf{v}(\mathbf{x},t)$, the equation for $S_{2n}(r)$ is given by [1,2,9]

$$-\frac{2}{r^{d-1}} \frac{\partial}{\partial r} \left(r^{d-1} \eta(r) \frac{\partial S_{2n}(r)}{\partial r} \right) = J_{2n}(r), \quad (2)$$

where d is the space dimensionality and $\eta(r)$ is the two-particle eddy-diffusivity scalar with the form $\eta(r) = \eta(L)(r/L)^{(2-\zeta_2)}$, L denoting the integral scale. The molecular-diffusion terms $J_{2n}(r)$ in Eq. (2) are

$$J_{2n}(r) = 2n\kappa \langle T_r^{2n-1} (\nabla_x^2 + \nabla_{x'}^2) T_r \rangle. \quad (3)$$

The existence of a power law for $S_{2n}(r)$ requires [2,9]

$$J_{2n}(r) = nC_{2n} J_2 \frac{S_{2n}(r)}{S_2(r)}, \quad (4)$$

where C_{2n} are dimensionless constants that must be determined. Equation (4) is also a consequence of the fusion rule proposed by L'vov and Procaccia [13–15]. From Eqs. (2) and (4), the scaling exponents ζ_{2n} can be evaluated [2]:

$$\zeta_{2n} = \frac{1}{2} [\sqrt{4nC_{2n}d\zeta_2 + (d - \zeta_2)^2} - (d - \zeta_2)]. \quad (5)$$

Based upon the linear ansatz $\langle (\nabla_x^2 + \nabla_{x'}^2) T_r | T_r \rangle \propto T_r$, the closure $C_{2n} = 1$ was proposed by Kraichnan [2]. The scaling exponents ζ_{2n} of the Kraichnan conjecture are $\zeta_{2n} = \frac{1}{2} [\sqrt{4nd\zeta_2 + (d - \zeta_2)^2} - (d - \zeta_2)]$.

Numerical simulations and experimental data [3,9,15] indicate that the linear ansatz may be approximately correct, but a small departure from the linear ansatz (especially for small values of T_r) can create large changes in the anomalous exponents. Therefore, there have been many studies of the scaling of a passive scalar using alternative methods [4–8]. Ching formulated the RSH as

$$T_r^2 = \theta^2 \frac{L^2}{\eta(L)} \left(\frac{r}{L} \right)^{\zeta_2} \chi_r, \quad (6)$$

where χ_r is the locally averaged scalar dissipation rate defined by $\chi_r = (1/V_r) \int_{B_r} \kappa |\nabla T(\mathbf{y})|^2 d\mathbf{y}$ with B_r being a d -dimensional ball centered at \mathbf{x} with radius r and volume V_r and θ a dimensionless random variable independent of r and statistically independent of $\chi(r)$. Using the formulation (6), Ching [12] evaluated the molecular-diffusion terms $J_{2n}(r)$ and showed that the scaling exponents ζ_{2n} satisfy

$$\zeta_{2n} = \frac{1}{2} \{ \sqrt{[d - \zeta_2 - g(n)\zeta_2]^2 + 4ng(n)\zeta_2(d - \zeta_2)} - [d - \zeta_2 - g(n)\zeta_2] \}, \quad (7)$$

where $g(n) = (2n-1) \langle X^{2n-2} \rangle / \langle X^{2n} \rangle$, in which $X = \theta / \sqrt{\langle \theta^2 \rangle}$. From Eq. (7), we can find all the integer moments of the random variable X if the scaling exponents ζ_{2n} are already known.

The statistical distribution of X can be determined uniquely from its integer-moments $\langle X^n \rangle$ if the power series $\sum_{n=1}^{\infty} [\langle X^{2n} \rangle \lambda^n / (2n)!]$ converges in some interval $|\lambda| < \lambda_0$, where $\lambda_0 > 0$ is the radius of convergence [16]. This sufficient condition is equivalent to $\lim_{n \rightarrow \infty} \langle X^{2n} \rangle / [2n(2n-1) \langle X^{2n-2} \rangle] = \lim_{n \rightarrow \infty} 1/[2ng(n)] < 1/\lambda_0 < \infty$. From Eq. (7), $\zeta_{2n} \geq \zeta_2$ leads to $g(n) \geq d/[nd + (1-n)\zeta_2]$. Thus, for the scaling of a passive scalar, the sufficient condition is satisfied naturally, so that the statistical distribution of X can be determined uniquely from its moments.

In the following, we give analytical solutions for the probability density function (PDF) $P(X)$ for the scaling at the two limits of n -independent scaling $\zeta_{2n} = \zeta_2$ and regular scaling $\zeta_{2n} = n\zeta_2$, and numerical calculations for the scaling of the Kraichnan conjecture.

For n -independent scaling $\zeta_{2n} = \zeta_2$, from $g(n) = d/[nd + (1-n)\zeta_2]$, we have

$$\int X^{2n} P(X) dX = \int (2n-1) \times \left[\left(1 - \frac{\zeta_2}{d} \right) n + \frac{\zeta_2}{d} \right] X^{2n-2} P(X) dX. \quad (8)$$

Using the method of Sinai and Yakhot [17], from Eq. (8) we can obtain

$$\int \left[\frac{1}{2} \left(1 - \frac{\zeta_2}{d} \right) X \frac{d^2 P}{dX^2} - \frac{\zeta_2}{d} \frac{dP}{dX} - XP \right] X^{2n-1} dX = 0. \quad (9)$$

Since $P(X)$ is an even function of X , we have

$$\frac{d^2 P}{dX^2} - \frac{2\nu}{X} \frac{dP}{dX} - c^2 P = 0, \quad (10)$$

where $\nu = \zeta_2 / (d - \zeta_2)$ and $c = \sqrt{2d / (d - \zeta_2)}$. Replacing X and $P(X)$ by $z = cX$ and $P = (c^{-1}z)^{\nu+1/2} F(z)$, we can turn Eq. (10) into the Bessel modified equation $d^2 F/dz^2 + (1/z) dF/dz - [1 + (\nu+1/2)^2/z^2] F = 0$. Noting that $\lim_{X \rightarrow \pm\infty} P(X) = 0$, we have $P(X) = A |X|^{\nu+1/2} K_{\nu+1/2}(c|X|)$ where $K_{\nu+1/2}(c|X|)$ is a modified Bessel function of the second kind of order $\nu+1/2$ and A is a constant. Determining the constant A from the condition $\int_{-\infty}^{+\infty} P(X) dX = 1$, the solution for Eq. (10) is

$$\begin{aligned} P(X) &= \frac{c^{\nu+3/2}}{2^{\nu+1/2} \Gamma(\nu+1) \Gamma\left(\frac{1}{2}\right)} K_{\nu+1/2}(c|X|) \\ &= \frac{c^{2\nu+2}}{\pi} |X|^{2\nu+1} \int_0^{\infty} \frac{\cos(t)}{(t^2 + c^2 X^2)^{\nu+1}} dt \\ &= \frac{c}{\pi} \int_0^{\infty} \frac{\cos(c|X|t)}{(1+t^2)^{\nu+1}} dt. \end{aligned} \quad (11)$$

From the analytical solution (11), we can obtain the following results.

(1) For $\zeta_2 \rightarrow 0$ or $d \rightarrow \infty$, we have $\nu \rightarrow 0$. Thus, the statistical distribution of X tends to the exponential distribution

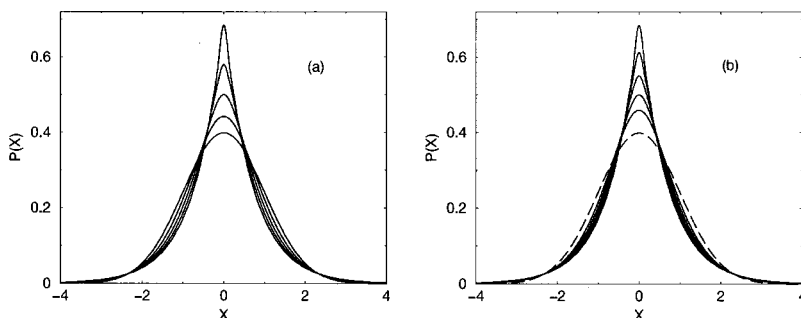


FIG. 1. Plot of the probability density function $P(X)$ as a function of $X = \theta / \sqrt{\langle \theta^2 \rangle}$ for n -independent scaling $\zeta_{2n} = \zeta_2$. (a) $d=2$ for different $\zeta_2 = 0.0$ (the exponential distribution), 0.5, 1.0, 1.5, 2.0 (the Gaussian distribution) (from top to bottom), and (b) $d=3$ for different $\zeta_2 = 0.0$ (the exponential distribution), 0.5, 1.0, 1.5, 2.0, and the Gaussian distribution (dashed line) (from top to bottom).

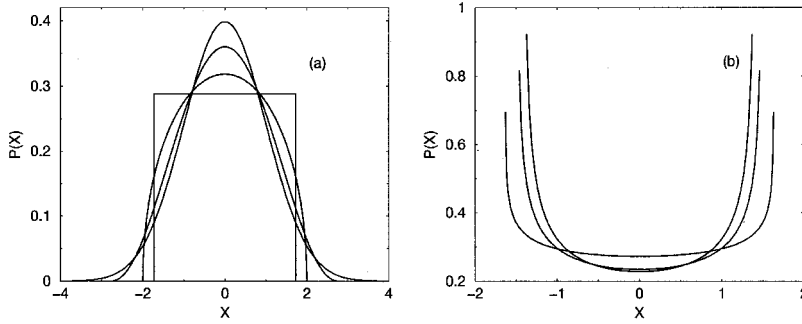


FIG. 2. Plot of the probability density function $P(X)$ as a function of $X = \theta/\sqrt{\langle \theta^2 \rangle}$ for regular scaling $\zeta_{2n} = n\zeta_2$ in $d=2$ for (a) $\zeta_2 = 0.0$ (the Gaussian distribution), 0.5, 1.0, 4/3 (the uniform distribution in $[-\sqrt{3}, \sqrt{3}]$) (from top to bottom) and (b) $\zeta_2 = 1.5, 1.8, 2.0$ (from bottom to top).

$$\lim_{\zeta_2 \rightarrow 0, d \rightarrow \infty} P(X) = \frac{1}{\sqrt{2}} e^{-\sqrt{2}X}.$$

(2) In $d=2$, the functions $g(n)$ tend to 1 when $\zeta_2 \rightarrow 2$. In this case, the statistical distribution of X will tend to the Gaussian distribution

$$\lim_{\zeta_2 \rightarrow 2} P(X, d=2) = \frac{1}{\sqrt{2\pi}} e^{-X^2/2}.$$

Figures 1(a) and 1(b) show the probability density function $P(X)$ in $d=2$ and $d=3$ for different ζ_2 .

For regular scaling $\zeta_{2n} = n\zeta_2$, the relation $g(n) = n\zeta_2/d + (1 - \zeta_2/d)$ gives

$$\begin{aligned} & \int \left[\frac{\zeta_2}{d}n + \left(1 - \frac{\zeta_2}{d}\right) \right] X^{2n} P(X) dX \\ &= \int (2n-1) X^{2n-2} P(X) dX. \end{aligned} \quad (12)$$

This implies

$$\left(1 - \frac{\zeta_2}{2d} X^2\right) \frac{dP}{dX} = \left(\frac{3\zeta_2}{2d} - 1\right) X P. \quad (13)$$

Since $P(X)$ is a non-negative function with the condition $\int_{-\infty}^{+\infty} P(X) = 1$, $P(X)$ is concentrated in a finite interval $[-\sqrt{2d/\zeta_2}, \sqrt{2d/\zeta_2}]$:

$$P(X) = B \left(1 - \frac{\zeta_2}{2d} X^2\right)^{d/\zeta_2 - 3/2}, \quad (14)$$

where $B = [\sqrt{2d/\zeta_2} \Gamma(1/2) \Gamma(d/\zeta_2 - 1/2) / \Gamma(d/\zeta_2)]^{-1}$. So we have the following results.

(1) For $\zeta_2 \rightarrow 0$ or $d \rightarrow \infty$, $g(n) \rightarrow 1$. The statistical distribution of X tends to a Gaussian distribution.

(2) In $d=2$, when $\zeta_2 = 4/3$, the random variable X has a uniform distribution in $[-\sqrt{3}, \sqrt{3}]$. It is interesting that the probability density function $P(X)$ is a convex function for $\zeta_2 < 4/3$ and a concave function for $\zeta_2 > 4/3$ [see Figs. 2(a) and 2(b)]. This appears to be a phase transition and the critical point is $\zeta_2 = 4/3$. In $d=3$, the situation is different. The probability density function $P(X)$ is a convex function for all $0 < \zeta_2 < 2$ and tends to a constant on its support $[-\sqrt{3}, \sqrt{3}]$ when $\zeta_2 \rightarrow 2$ [see Fig. 3].

It is surprising that similar distribution to (14) has been found in the investigation of the RSH of a turbulent velocity field [11] and for the statistical quasistationary PDF of the passive scalar $X = (T - \langle T \rangle) / \langle (T - \langle T \rangle)^2 \rangle^{1/2}$ in turbulence [18].

For the scaling of the Kraichnan conjecture $\zeta_{2n} = \frac{1}{2} [\sqrt{4nd\zeta_2 + (d - \zeta_2)^2} - (d - \zeta_2)]$, we can calculate the moments $\langle X^{2n} \rangle$ for all positive integers n from $g(n) = 1.0 / (1.0 - \zeta_2/d + \zeta_{2n}/nd)$. The characteristic function $f(t)$ of the probability density function $P(X)$ is

$$f(t) = 1 + \sum_{n=1}^{\infty} \frac{(-1)^n \langle X^{2n} \rangle}{(2n)!} t^{2n}.$$

Noting that $f(t)$ is an even function of t , we have

$$P(X) = \frac{1}{2\pi} \int_{-\infty}^{+\infty} f(t) e^{-ixt} dt = \frac{1}{\pi} \int_0^{+\infty} f(t) \cos(xt) dt.$$

Using the technique for numerical integration of an oscillating function [19], we obtain the numerical results for the probability density function $P(X)$ in $d=2$ and $d=3$ [see Figs. 4(a) and 4(b)]. Based upon the numerical calculations, we have the following results.

(1) For $\zeta_2 \rightarrow 0$ or $d \rightarrow \infty$, the statistical distribution of X tends to the Gaussian distribution.

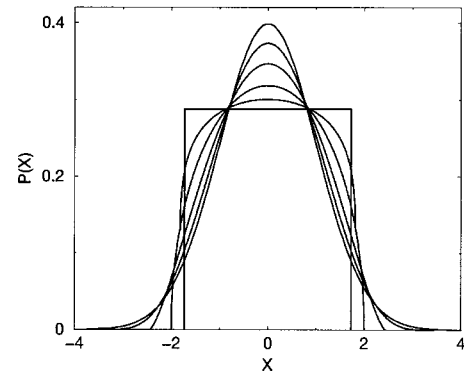


FIG. 3. Plot of the probability density function $P(X)$ as a function of $X = \theta/\sqrt{\langle \theta^2 \rangle}$ for regular scaling $\zeta_{2n} = n\zeta_2$ in $d=3$ for different $\zeta_2 = 0.0, 0.5, 1.0, 1.5, 1.8, 2.0$ (from top to bottom) where $\zeta_2 = 0.0$ and 2.0 correspond to the Gaussian distribution and the uniform distribution in $[-\sqrt{3}, \sqrt{3}]$, respectively.

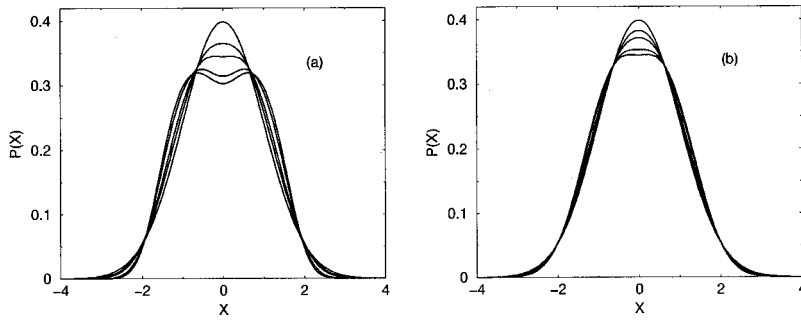


FIG. 4. Plot of the probability density function $P(X)$ as a function of $X = \theta/\sqrt{\langle \theta^2 \rangle}$ for the Kraichnan conjecture for different $\zeta_2 = 0.0$ (the Gaussian distribution), 1.0, 4/3, 1.8, 2.0 (from top to bottom) in (a) $d=2$ and (b) $d=3$.

(2) It is interesting that the probability density function $P(X)$ has a saddle shape with two peaks in $d=2$ when $\zeta_2 > 1.33$. In $d=3$, it has only one peak at $X=0$ for all $0 < \zeta_2 < 2$ and changes very slowly with the value of X in the neighborhood of $X=0$ when $\zeta_2 \rightarrow 2$.

Therefore, we suggest that the statistical distribution of X for the scaling of the Kraichnan conjecture has some similar properties to the regular scaling $\zeta_{2n} = n\zeta_2$ in the neighborhood of $X=0$ and also some similar properties to

n -independent scaling $\zeta_{2n} = \zeta_2$, e.g., $P(X)$ seems to be concentrated in an infinite interval $(-\infty, \infty)$.

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